

Dependency/Competition in Corn Yield: A Proxy for Land Use Change

October 2009

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Setup

We investigate the effect of biofuels on land use change via U.S. corn production data.

- ▶ Agricultural models are used to estimate the effect of biofuel production on crop production $\Rightarrow$ land use change
- ▶ Statistical determination of the influence of biofuel production is non-standard $\Rightarrow$ total crop use is always the sum of the different uses
- ▶ These (induced) distributions are *Compositional Distributions* $\Rightarrow$ Methodology for Dependency/Competition among Compositional Distributions

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Outline

Context

Methodology

Illustrations

Comments

Energy Independence and Security Act

- ▶ EISA-2007 mandates an increase in ethanol production to 36 billion gallons per year by 2022.
- ▶ Ethanol production has increased more than 5000 percent since 1980
- ▶ At the same time, the total U.S. corn yield has less than doubled

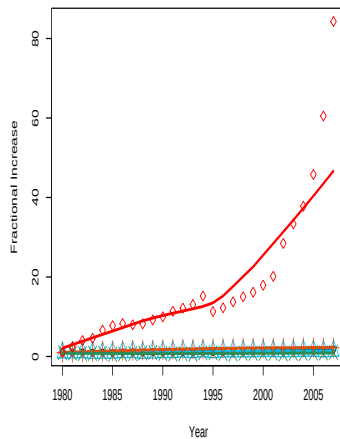
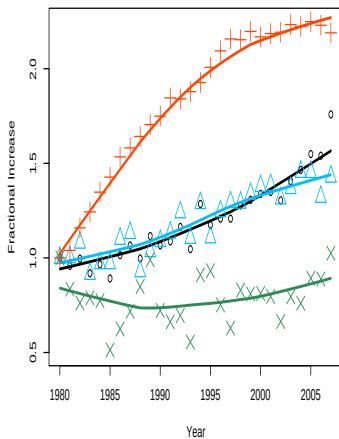
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Fractional Increase



Environmental Impacts

- ▶ **Net energy budget**
- ▶ Competition with corn-based commodities
- ▶ Greenhouse emissions
- ▶ $\Rightarrow$ Competition within distribution of Corn Yield constituents

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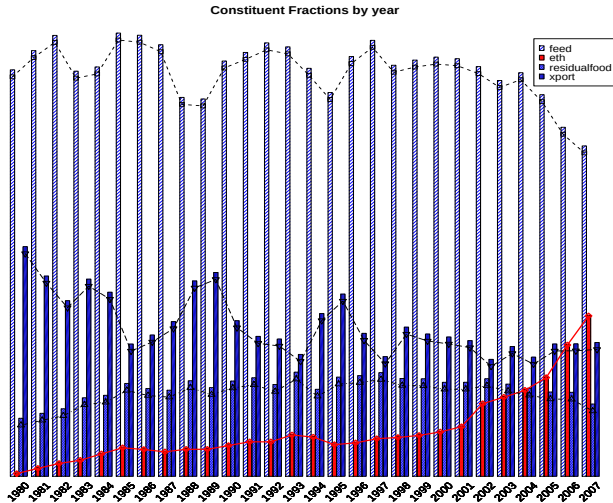
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Constituents of Corn Yield



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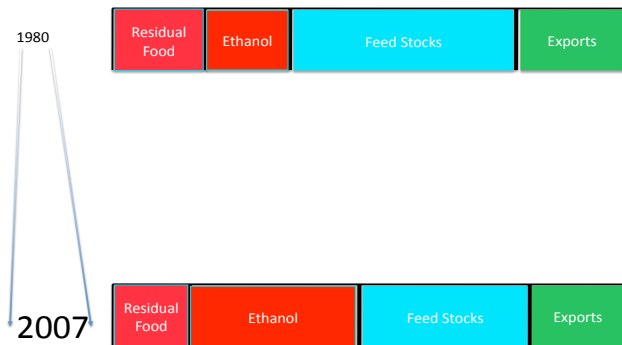
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Compositional Distribution



Aitchison Notation

Let

$$\mathbf{x} = (x_1, \dots, x_k) \quad (1)$$

be a basis or open vector of positive quantities

In this example

$$\mathbf{x} = (x_{eth}, x_{rfood}, x_{feed}, x_{xport})$$

(in bushels) corn of: ethanol production, residual food stock, feed stock, and exports.

Aitchison Notation

Let

$$y_j = x_j / \sum_j^k x_j \quad (2)$$

$\mathbf{y} = (y_1, \dots, y_k)$ the vector of fractions.

Aitchison defines $y_{k+1} = 1 - \sum_j^k y_j$;

Here $\sum_j^k y_j = 1$

Aitchison Notation

A (log-ratio) transformation sets

$$v_{j \atop m} = \log\left(\frac{y_j}{y_m}\right) = \log y_j - \log y_m \quad (3)$$

in a slight modification of Aitchison's notation
(where $v_j = \log(y_j/y_{k+1})$).

Modifying Aitchison Notation

- ▶ The total is fixed and known $\Rightarrow$ the residual is $y_{k+1}=0$ and Aitchison's v_j is undefined
- ▶ In the original notation v_j is the log of the relative fraction of constituent j to the residual component of the basis
- ▶ $v_{\cdot, m}$, maps $\mathbb{S}^{m,k} = \{\mathbf{y}_{-m}, y_m\}$ to $\mathbb{R}^{k-|m|}$

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Why Transform?

(Natural) Dirichlet model for $(y_1, \dots, y_k)$; $\sum_j y_j = 1$; $y_j > 0 \forall j$ is:

$$dF(\mathbf{y}) \propto (1 - \sum_j y_j)^{\alpha_{k+1}-1} \cdot \prod_j y_j^{\alpha_j-1} \quad (4)$$

with parameters $\alpha = (\alpha_1, \dots, \alpha_{k+1})$

Insufficient for non-*neutral* proportions

Why Transform?

(Generalization) Liouville distribution is:

$$dF \propto h\left(\sum_j y_j\right) \prod y_j^{\alpha_j-1} \quad (5)$$

with $\alpha_j > 0$ (as before) and g some function.

Note that when $h(t) = 1 - t$ the Liouville distribution is the special case Dirichlet distribution with $\alpha_{k+1} = 1$

Thus h is an additional parameter of interest for estimation

Aitchison Approach

- ▶ Aitchison fits a log-normal distribution on $\mathbf{v}$
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- Under a composition

$$\Sigma_{\mathbf{v}} \propto \text{diag}(\omega_1, \dots, \omega_k) + \omega_{k+1}, \quad (6)$$

- $\Sigma_{\mathbf{v}}$ is constrained to the positive orthant and proportional to the units of the residual component

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Aitchison Approach

- ▶ Aitchison uses a likelihood ratio test
- ▶ Test statistic is iteratively estimated due to the constraints on the support of the parameter space
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Our Approach

Multivariate Version of KS distance:

$$D_{n,k} = \sup_{\mathbf{t}} |F_n(\mathbf{t}) - F(\mathbf{t})| \quad (7)$$

For $\mathbf{t} = (t_1, t_2, \dots)$ the distance is a probability measure on Kendall's distributions...chi-square convergence does not hold (Nelsen 2003).

Our Approach

A Kolmogorov-Smirnov statistic (distance) for multivariate independence can be written:

$$D_{n,k}^{\Pi} = \sup_{\mathbf{t}} |F_n(\mathbf{t}) - \prod_j F_j(t_j)|. \quad (8)$$

Under independence the distance converges to zero (via Glivenko-Cantelli), but distributionally for $k \geq 2$?

Our Approach

- ▶ Let $\mathbf{u} = (u_1, \dots, u_k)$, where each $u_j = F_j(v_j)$
- ▶ Let the joint distribution for $\mathbf{v}$ be $F(\mathbf{v})$
- ▶ The *copula* for $\mathbf{u}$ is

$$C(\mathbf{u}) = F(F_1(v_1), \dots, F_k(v_k)) \quad (9)$$

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With $C_n(\cdot)$ a multivariate version of the *empirical copula*:

$$C_n(\mathbf{u}) = \frac{\#\{\mathbf{t} \mid t_1 \leq F_1^{-1}(u_1), \dots, t_k \leq F_k^{-1}(u_k)\}}{n} \quad (10)$$

Our Approach

- ▶ Fit a *Dirichlet* distribution (i.e. estimate $\hat{\alpha} = (\hat{\alpha}_1, \dots, \hat{\alpha}_k)$ for $\alpha = (\alpha_1, \dots, \alpha_k)$) to the composition data $\mathbf{y}$.
- ▶ Generate T *Dirichlet* replicates, parameter $\hat{\alpha}$, each of dimension $n \times k$: $(\mathbf{y}^{\hat{\alpha},1}, \dots, \mathbf{y}^{\hat{\alpha},T})$.
- ▶ Compute $m = 1 \dots k$ versions of Aitchison's log-ratios on the replicates: $\mathbf{v}_m^{\hat{\alpha},1}, \dots, \mathbf{v}_m^{\hat{\alpha},T}$
- ▶ For $m = 1 \dots k$ compute $D_{n,k}^{\Pi,1}, \dots, D_{n,k}^{\Pi,T}$ of

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- ▶ m versions of $D_{n,k}^{\Pi,1}, \dots, D_{n,k}^{\Pi,T}$ are proxies for tests of *complete subcompositional independence*...
- ▶ Calculating on the log-ratios ($\mathbf{v}$) of the replicates, and not the *Dirichlet* draws picks each of m components to serve as 'residual' (via the basis $\mathbf{x}$ or composition $\mathbf{y}$) without requiring m estimates of α and m -fold random draws.

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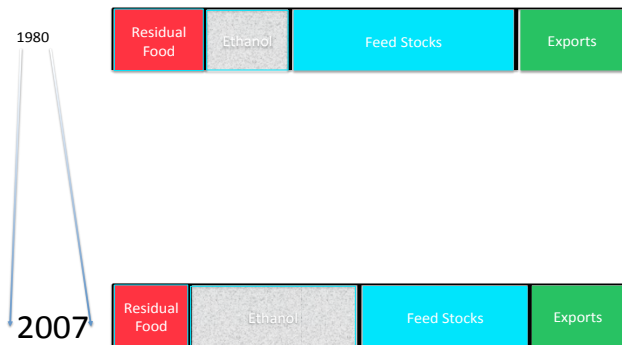
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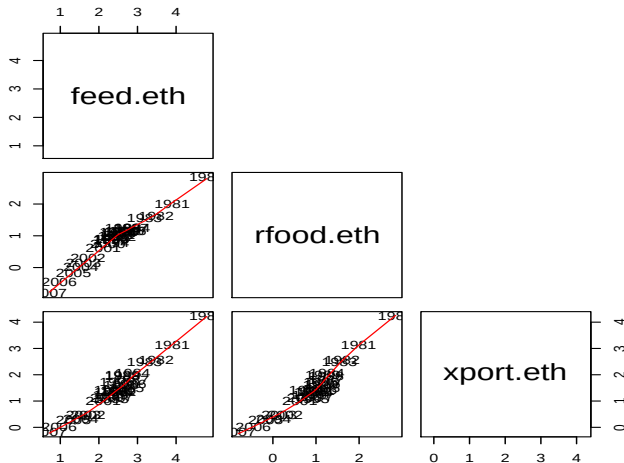
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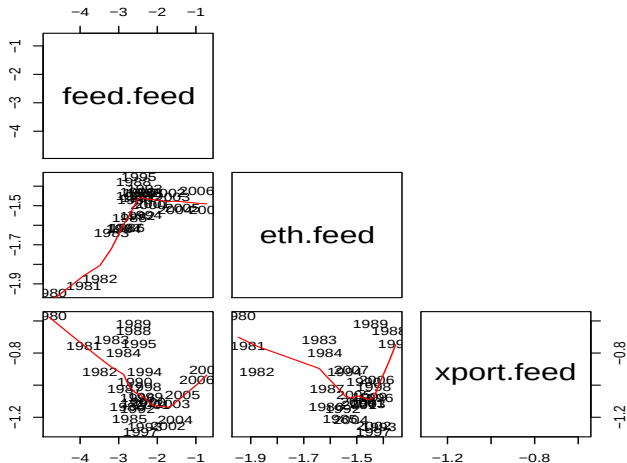
Compositional Distribution



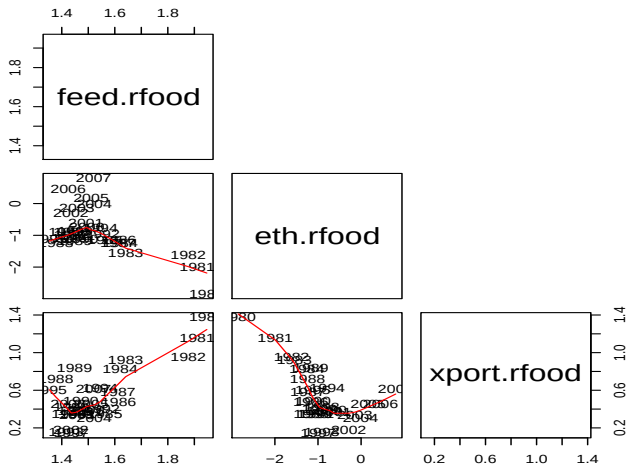
W.R.T. Ethanol



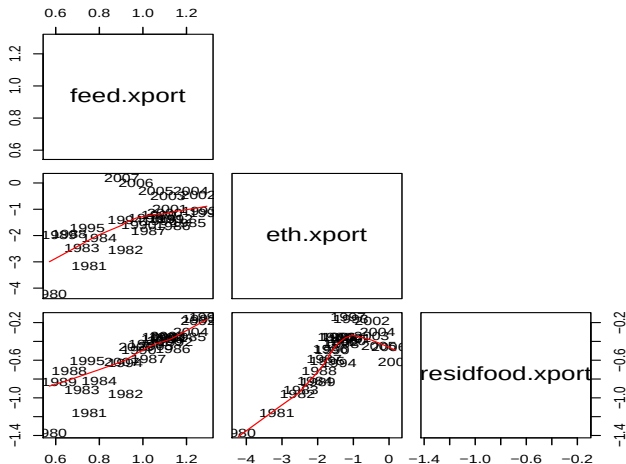
W.R.T. Feed



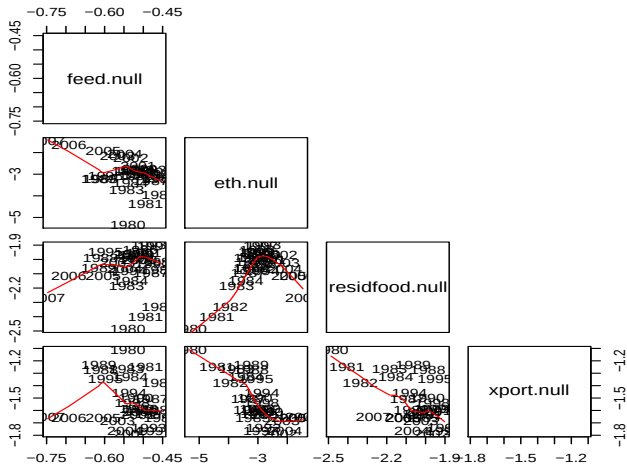
W.R.T. Food



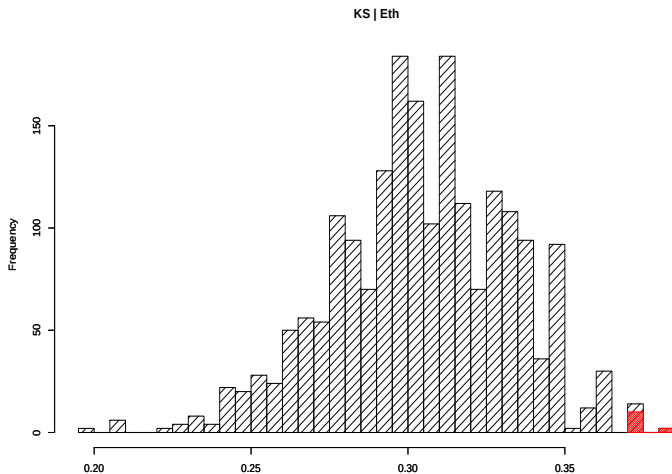
W.R.T. Export



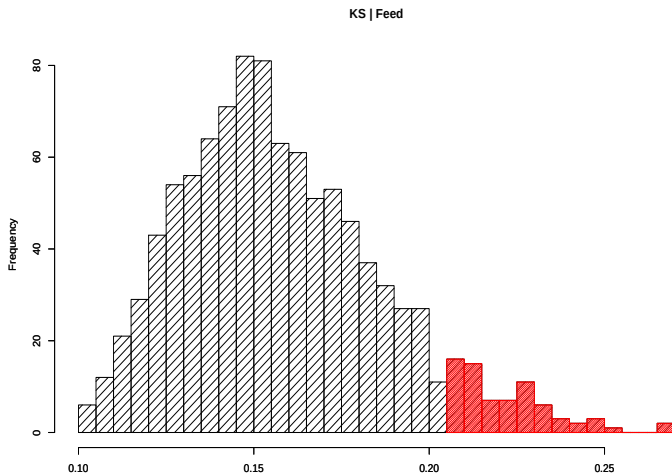
W.R.T. Null



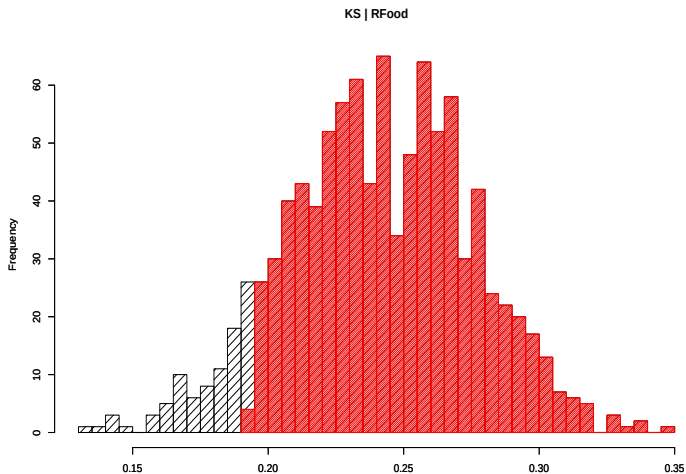
W.R.T. Ethanol



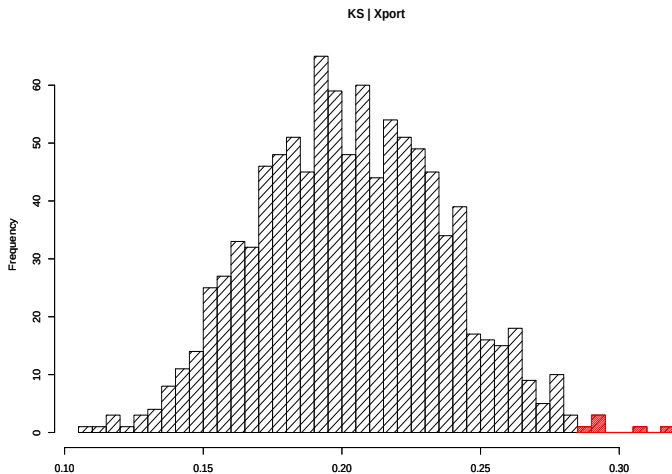
W.R.T. Feed



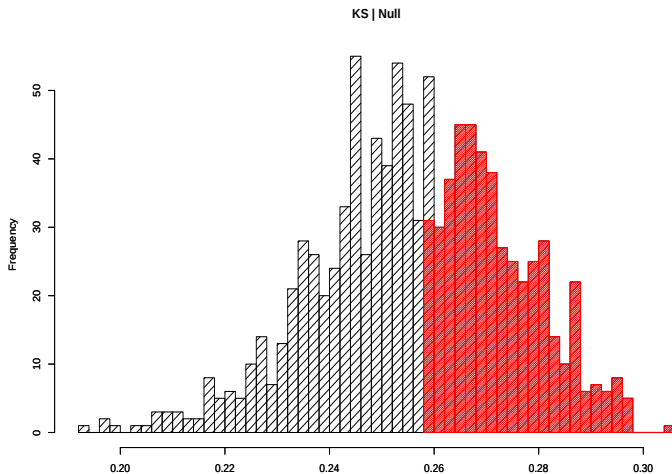
W.R.T. Residual Food



W.R.T. Export



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- ▶ Empirical Prob Integral Transform : : Order statistics $\Rightarrow$ Invariant to increasing (log-ratio) transform.
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