Straightforward (yet Novel) Methodology for Inequality: Conditional Lorenz Curves Duke University Conference on Social Determinants of Health Disparities August 2011

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Motivation

Constrained Sum Data

Inequality as a Measurement

- Partition Inequality
 - Group-wise
 - Contribution-wise

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Motivation

Constrained Sum Data

Inequality as a Measurement

- Partition Inequality
 - Group-wise
 - Contribution-wise
- Statistically Specify Inequality
 - As data
 - From some 'Random' process
 - for tests of significant differences

GOAL: Straightforward (Easy) Conditional/Groupwise Estimates of Inequality, with Probability Intervals

Just a little notation

Brief Notation

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- ▶ $\mathbb{1}_{[y_1 \leq y]} \leftarrow$ Indicator function. Say y = 5 and $y_1 = 3$, $y_2 = 7$ then $\mathbb{1}_{[y_1 \leq y]} = 1$ but $\mathbb{1}_{[y_2 \leq y]} = 0$

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The ecdf in this context is just the proportion of people with a less or equal amount y of the 'good'

US Income Data - ecdf



Empirical Distribution (Function) on CPI-U-RS Money Income, 2008

Figure: Graph of empirical cumulative distribution function (ecdf) of Money Income of Households — Consumer Price Index Research Series Using Current Methods, CPI-U-RS

US Income Data - L-curve



L-Curves on *binned* CPI-U-RS Money Income, 2008

Figure: Illustration of L-curves calculated on US Census CPI-U-RS money income in 2008.

The median household net worth for white Americans is \$113,149, and for blacks it's \$5,677. That's not a misprint or a misunderstanding; the median white household is 20 times richer than the median black household.

Figure: Powerful Words

Essentially all functions of ecdf

'Information' based Theil Index:

$$T = N^{-1} \sum_{i=1}^{N} r_i \log r_i = \sum_{j=1}^{m} \pi_j r_j \log_b r_j + \sum_{j=1}^{m} \pi_j r_j T_j$$
(2)
$$r_i = y_i / \overline{y},$$

$$\pi_j \leftarrow \text{relative size of group j,}$$

$$T_j \leftarrow \text{fix group j.}$$

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Essentially all functions of ecdf

'Mean Absolute Deviation' Gini Index:

$$G = \frac{\binom{n}{2}^{-1}}{2} \sum_{i < j} |y_i - y_j|$$
(3)

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 - properly a function of ecdf

Lorenz Curve

The beautiful Lorenz Curve

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also numbers between 0 and 1.

Just a little more notation

Brief Notation



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The Lorenz curve is just the sorted, cumulative list of 'good' shares by population proportion.

 $\mathsf{Lorenz} \to \mathsf{Gini}$

The Gini coefficient is a function of the Lorenz curve...

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$$G = \frac{\frac{1}{2} - \sum_{p=1/N}^{N} \frac{1}{N} L(p)}{1/2} = 1 - 2 \frac{1}{N} \sum_{p=1/N}^{N} L(p)$$
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...the scaled difference between the area under the observed Lorenz and equality $\label{eq:constraint}$

 $\mathsf{Lorenz} \to \mathsf{Gini}$

Lorenz Curves on CPI–U–RS Money Income, 2008


The trick is to see covariates as 'conditional information'

Aaberge et al [1] define pseudo-Lorenz regression curve as a function, in the presence of covariates \mathbf{x} for y, such that

$$E[\Lambda(p|\mathbf{x})] = L(p) \tag{7}$$

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e.g. that the conditional curves should 'sum' to the original curve

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$$L(\boldsymbol{p}) = \sum_{j=1}^{m} \pi_j \ \Lambda(\boldsymbol{p} | \mathbf{x} \in C_j)$$
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and setting

$$\Lambda(p|C_j) = \frac{\overline{y}_j}{\overline{y}} \cdot n_j \ L(F_j(F^{-1}(p))|C_j)$$
(9)

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guarantees that the overall Lorenz curve will be the weighted sum of conditional 'pseudo'-Lorenz curves. Just a little more notation

More Brief Notation

•
$$\pi_j = \frac{\overline{y}_j}{\overline{y}} \cdot n_j \leftarrow$$
 the proportional size of group j

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- L(F_j(F⁻¹(p))|C_j) ← the Lorenz curve of group j on the observed proportion of population in group j at the pth quantile of the overall distribution

In Layman's terms...

A simple algorithm

1. Sort all the data; Generate the *pth* quantiles of the unconditioned distribution. $\rightarrow F_N, F^{-1}(p)$

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- 2. Sort the data within each group; Generate the ecdf for each group (conditional distribution) at the *pth* quantiles, *of the original distribution*. $\rightarrow F_j(F^{-1}(p))$

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3. Join the *pth* proportions for each group F_j with the cumulative proportion of income at each group. $\rightarrow L(F_i(F^{-1}(p))|C_i)$

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- 3. Join the *pth* proportions for each group F_j with the cumulative proportion of income at each group. $\rightarrow L(F_j(F^{-1}(p))|C_j)$
- 4. Compute the contribution to the overall Lorenz curve, at each *pth* proportion. $\rightarrow \frac{\overline{y_j}}{\overline{v}} \cdot n_j L(F_j(F^{-1}(p))|C_j)$

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Simple Example



Illustrate the conditional lorenz curves for each group lnew1<-lorenz(g1); lnew2<-lorenz(g2); lnew3<-lorenz(g3)</pre>

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The function to compute the lorenz curve is sooooo easy
lorenz function(x)
y<-sort(x)
m<-mean(y); s<-sum(y)
l<-cumsum(y)/s
l
```

Simple Example



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Generally the 'resolution' can be set 'arbitrarily'. (but it's easy to set it at fewest group)

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```

```
groupsizesg<-c(4,4,4)/12 [1] 0.3333333 0.3333333 0.3333333
```

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```
for(pg in uppsg)
lpg<-c(lnew1[pg],lnew2[pg],lnew3[pg])
conditionallorenzedg[pg]<-
as.double(sum(meanratiosg*lpg*groupsizesg))
conditionallorenzedg
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And their contributions to the overall lorenz curve are

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Simple Example



And the Gini's are easy to compute ginioverall<-1-2*sum(conditionallorenzedg)/4


Example

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$$\frac{\partial L(p)}{\partial C} \bigg|_{C=c_j} \left[\sum_{j=1}^m \pi_j \ \frac{\overline{y}_j}{\overline{y}} \cdot n_j \ L(F_j(F^{-1}(p))|C_j) \right]$$
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But if we remember the definition of the derivative, and that the categorical covariate is 'singular', this is just

$$L(p)\Big|_{C_{-j}} - L(p)\Big|_C \tag{11}$$

Just the difference between the overall (conditionally defined) lorenz curve without and with the *jth* group.

Statistical significance

We can test for statistical significance using exploiting the duality between the Lorenz curve and the ecdf

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$$F_N(t) \sim N(F(t), F(t)[1 - F(t)])$$
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Then

$$L_N(p) \sim N(L(p), \frac{L(p)[1-L(p)]}{N})$$
(13)

and we can use normal confidence bounds (pointwise), or at least the Kolomorogov-Smirnov (KS) test for differences in distributions to test for significant effects. See [3].

We must be careful not to confuse data with the abstractions we use to analyze them.

-William James

Actually Done





Actually Done





References I

Rolf Aaberge, Steinar Bjerve, and Kjell Doksum.

Decomposition of rank-dependent measures of inequality by subgroups.

Metron - International Journal of Statistics, 63(3):493-503, 2005.

Kobi Abayomi and William Darity Jr. A friendly amendment to the theil index. Working paper, 2010.

Kobi Abayomi and Tracy Yandle.

A novel method of measuring consolidation, using conditional lorenz curves to examine itq consolidation in new zealand commercial fishing.

Marine Resources Research, 2011.